

# WHAT IS THE PROPER EVALUATION METHOD: some basic considerations

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**Scope:** provides best knowledge of observables

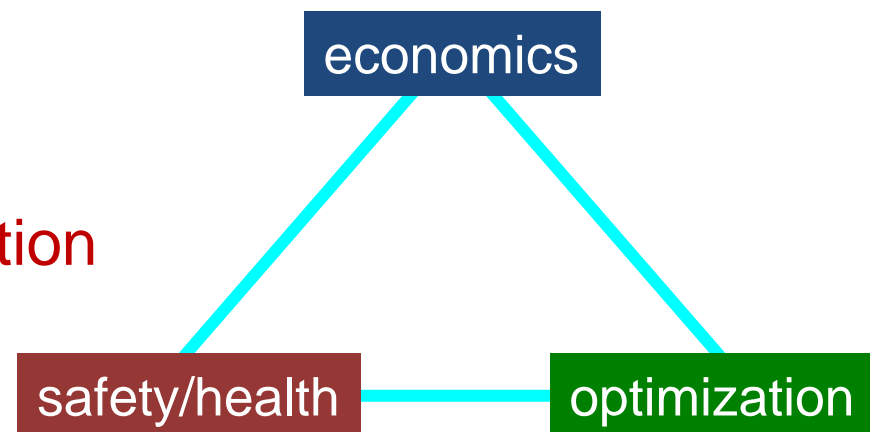
- consistent set of cross sections, spectra ...
- basic principles satisfied (unitarity, sum rules, ...)
- compatible with a-priori knowledge

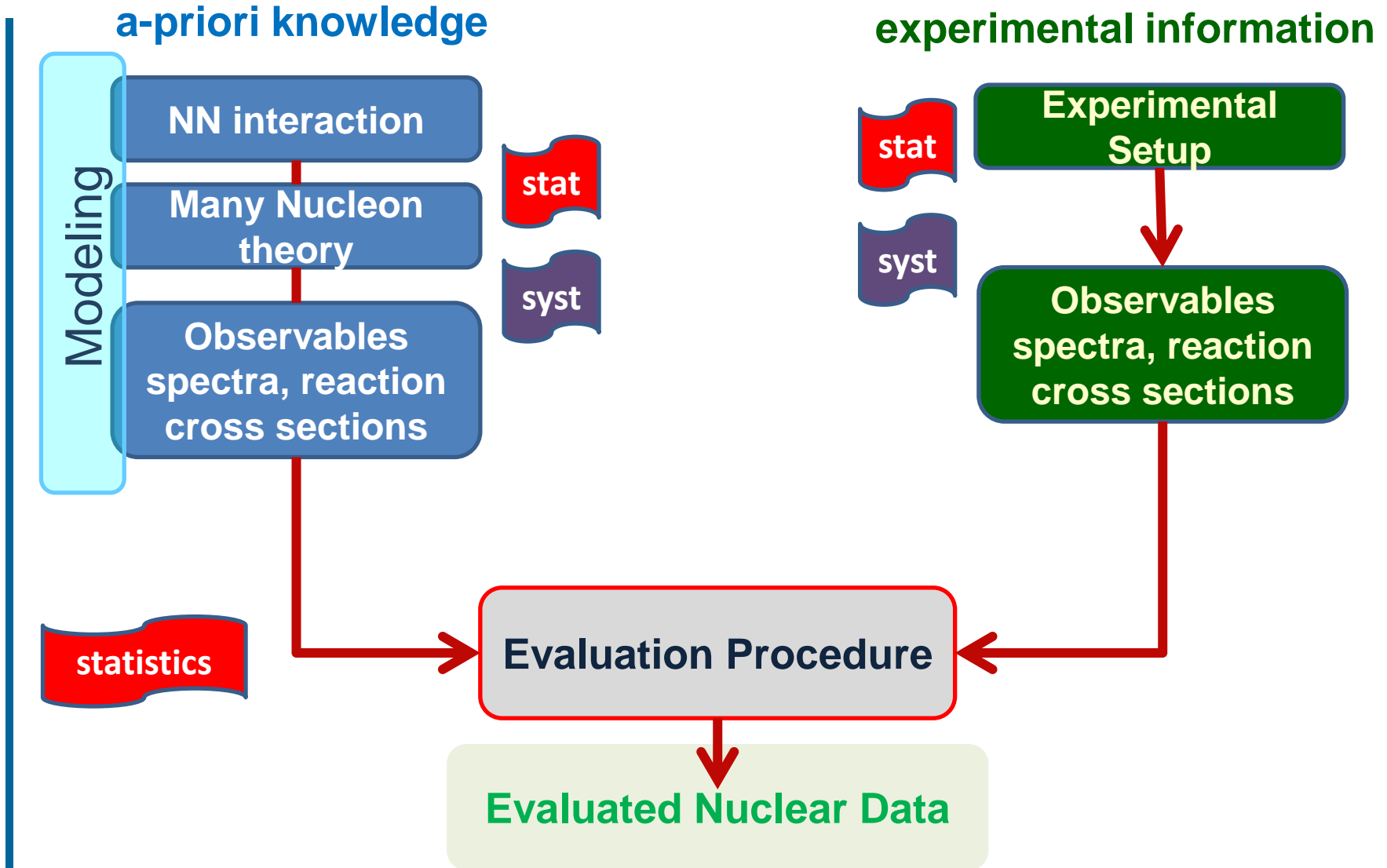
**Interest:** motivated by application

- design and construction of nuclear facilities (reactor, waste, fusion)
- accelerator applications (medicine, materials science, ADS)
- Safety issues and control systems (non-destructive testing, ... )

**Uncertainties:** driven by

**Requests:** reliability of evaluation





**Bayesian statistics:**

sum rule:  $p(\underline{x}|M) + p(\bar{\underline{x}}|M) = 1$

product rule:  $p(\underline{x}|\underline{\sigma} M) p(\underline{\sigma}|M) = p(\underline{\sigma}|\underline{x} M) p(\underline{x}|M)$

**Bayes Theorem (1763):**

$$p(\underline{x}|\underline{\sigma} M) = \frac{1}{p(\underline{\sigma}|M)} p(\underline{\sigma}|\underline{x} M) p(\underline{x}|M)$$

**updated probability distribution of the set of parameters**

**likelihood function, from experiment**

**PRIOR**

$$p(\underline{x}|\underline{\sigma} M)$$

.... probability distribution of parameters  $\underline{x}$  for a given model  $M$  and data  $\underline{\sigma}$

$$p(\underline{\sigma}|\underline{x} M)$$

.... probability distribution of data  $\underline{\sigma}$  for a model  $M$  with parameters  $\underline{x}$

$$p(\underline{x}|M)$$

.... probability for the occurrence of  $\underline{x}$  when  $M$  is true

# Nuclear Data Evaluation within Bayesian Statistics

**Prior:** obtained by variation of nuclear model parameters.  
(nuclear models for n-<sup>181</sup>Ta as given in TALYS 1.4 by default)

- physically reasonable boundaries defined
- homogenous distributions of parameters
- no correlation between parameters assumed

negative eigenvalues eliminated by regularization

**Experimental Data for n-<sup>181</sup>Ta:**

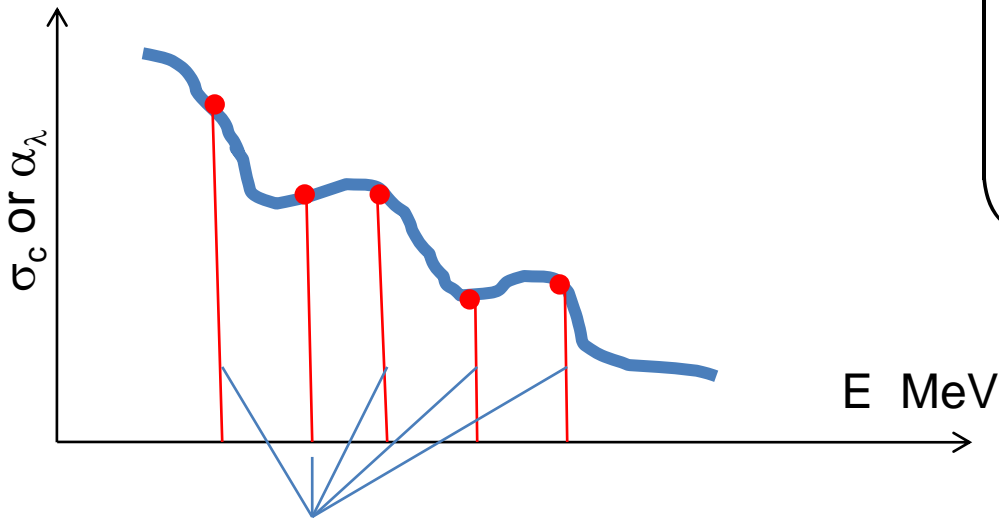
Total cross sections, elastic differential cross sections

**Evaluation:**

Linearized Update Procedure (assumption: normal distribution)

Use of cubic spline interpolation to represent  $\sigma_c(E)$  and  $\alpha_\lambda(E)$

Properties are conserved:  
especially sum rules

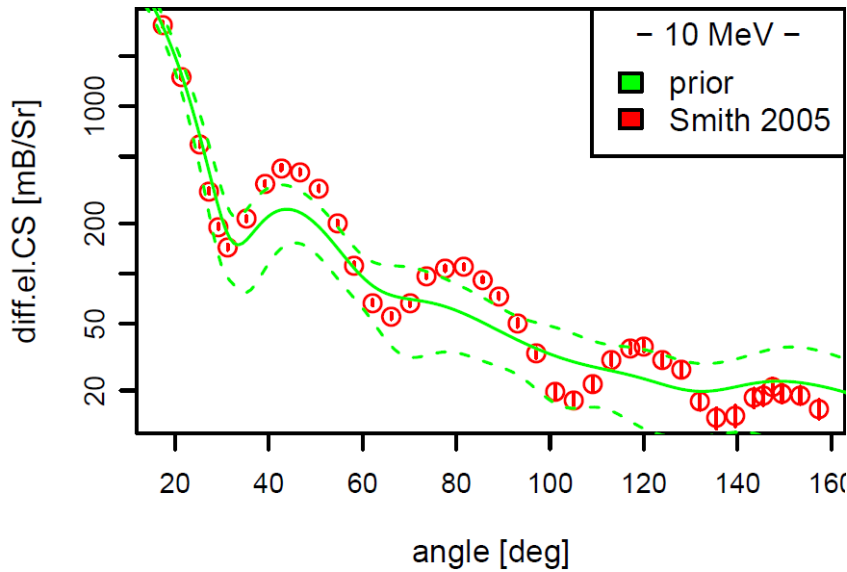


$$\begin{pmatrix} \sigma_c(E'_1) \\ \sigma_c(E'_2) \\ \vdots \\ \sigma_c(E'_J) \end{pmatrix} = \underline{\underline{S}} \begin{pmatrix} \sigma_c(E_1) \\ \sigma_c(E_2) \\ \vdots \\ \sigma_c(E_M) \end{pmatrix}$$

S is a  $J \times M$  matrix

values  $p_i$  for  $\sigma_c(E_i)$  or  $\alpha_\lambda(E_i)$  at the mesh points  $E_i$  are parameters of the cubic spline representation

# Evaluated differential elastic n-<sup>181</sup>Ta cross section at E=10 MeV

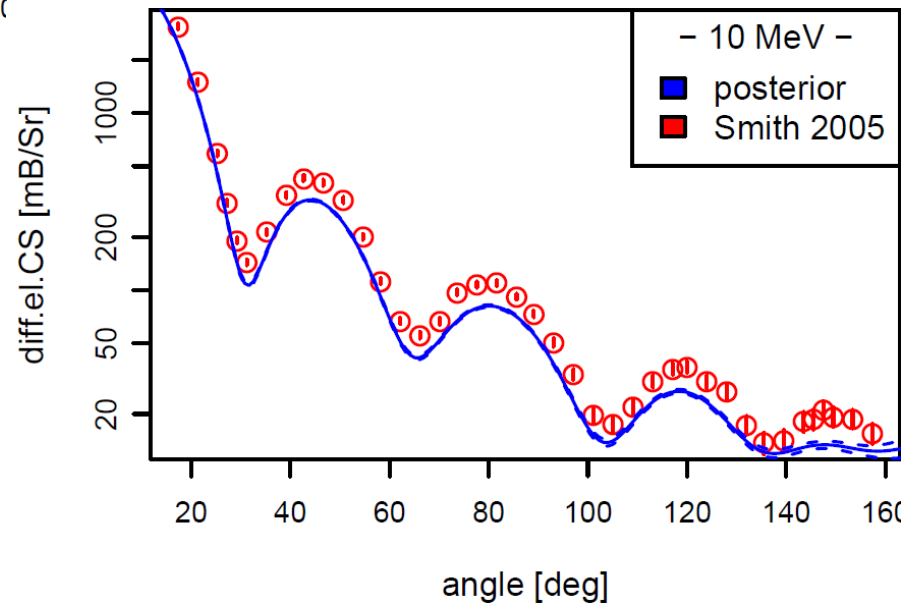


Problem:  
 The evaluated values are lower and the variances of the evaluated cross sections are too small.

Model defects negligible for total cross section

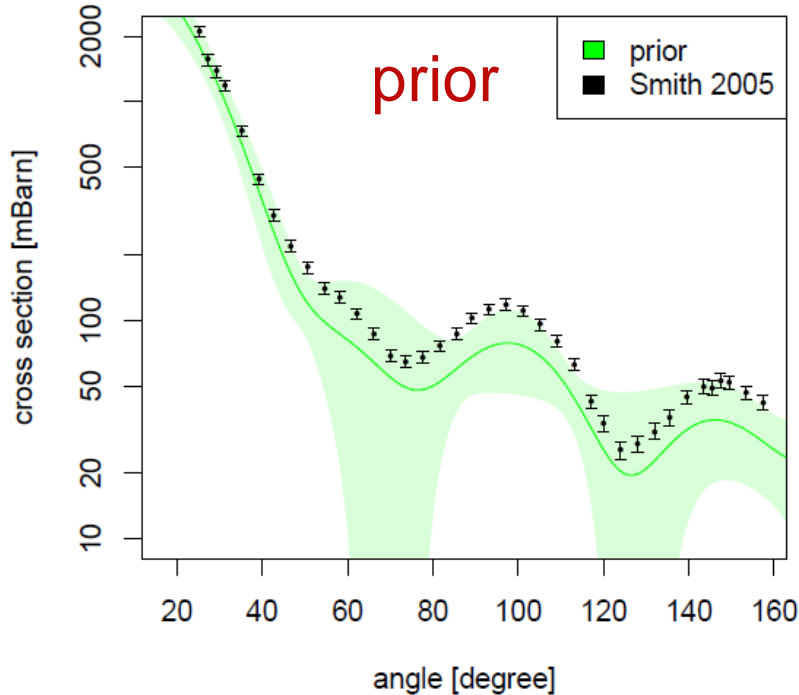
Question:

Is this specific for the energy or a general problem?



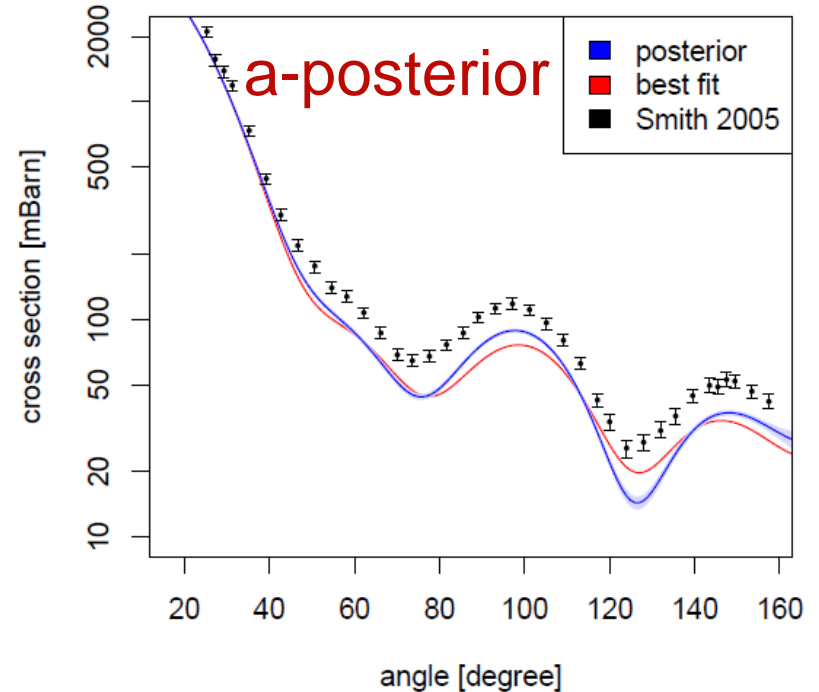
# Evaluated differential elastic n-<sup>181</sup>Ta cross section at E=4,51 MeV

elastic DA at 4.51 MeV

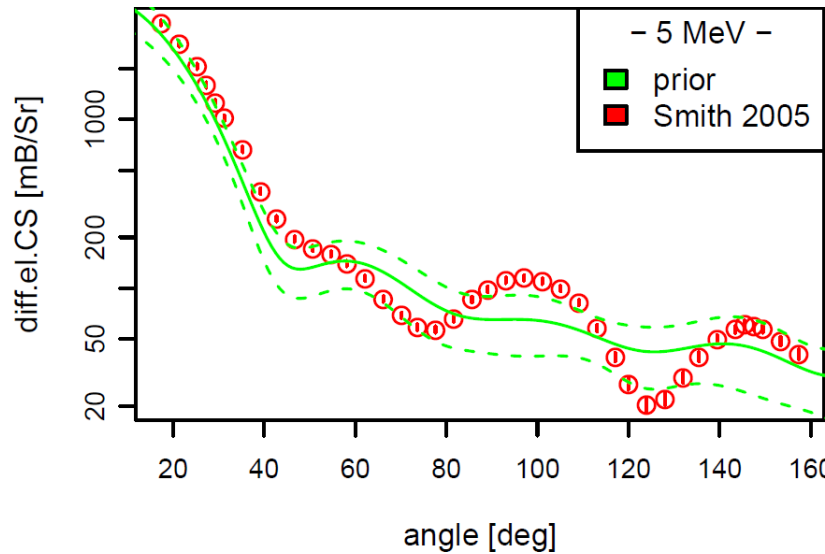


The effect is similar at all energies considered

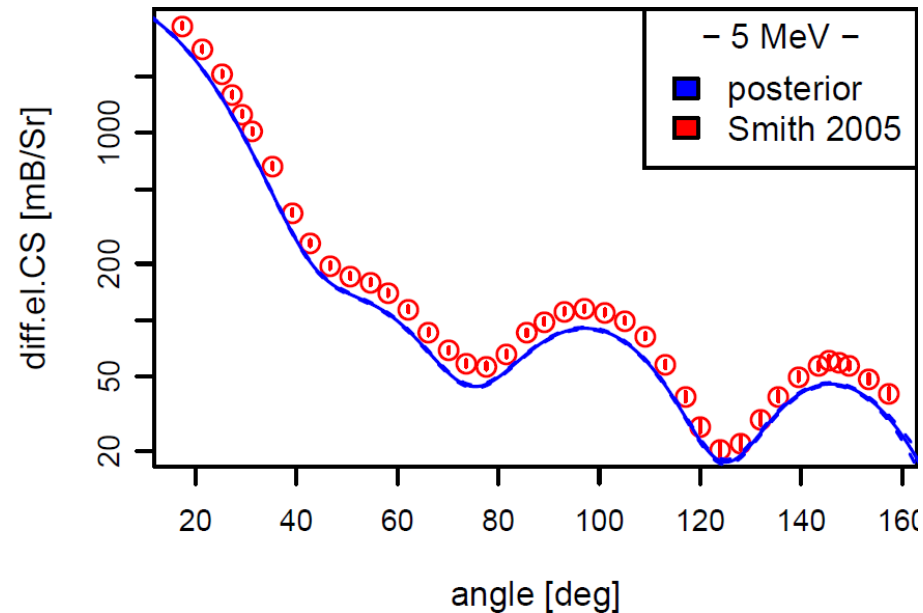
elastic DA at 4.51 MeV

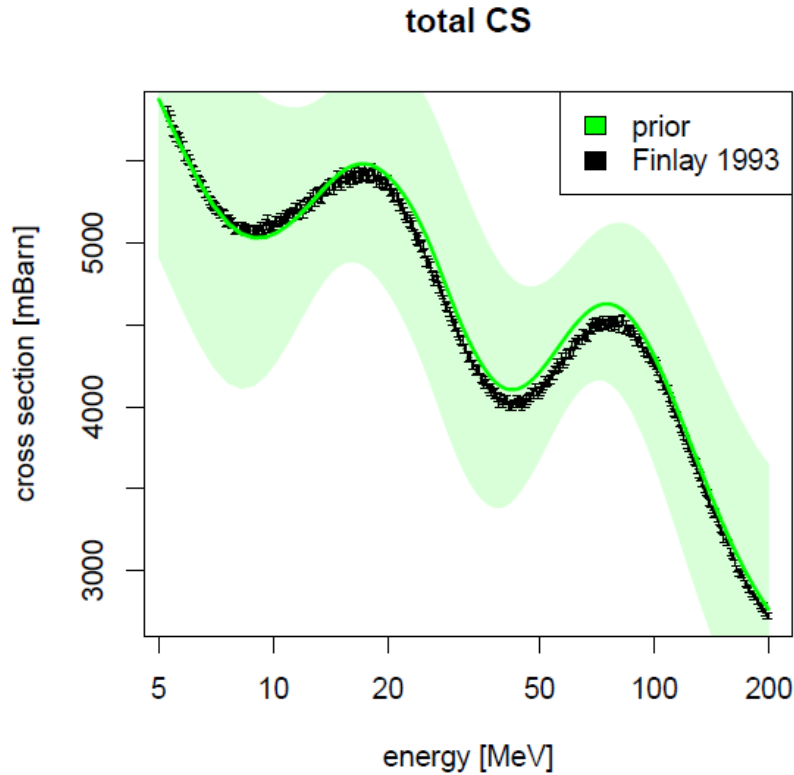




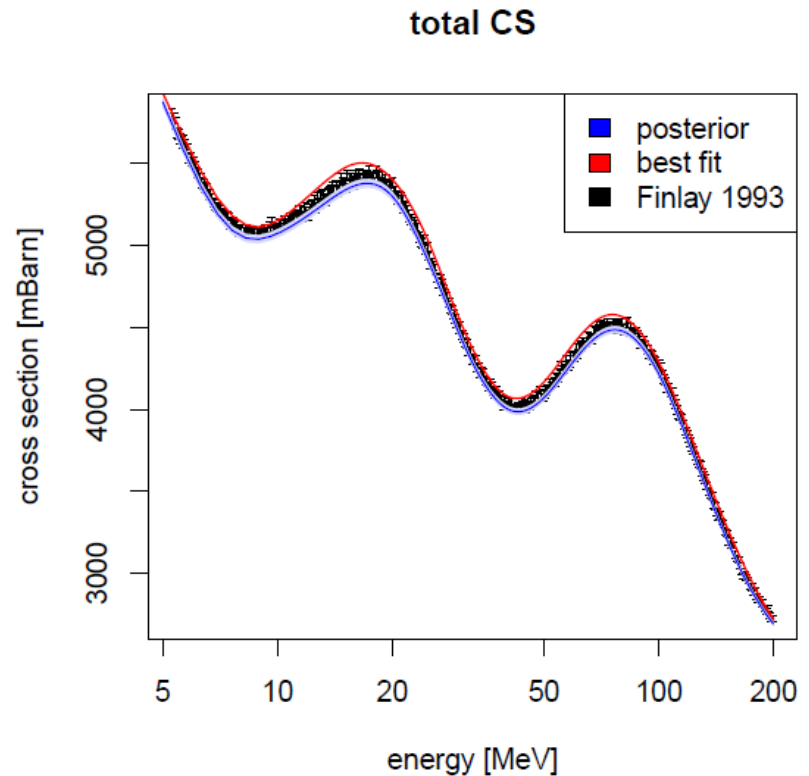


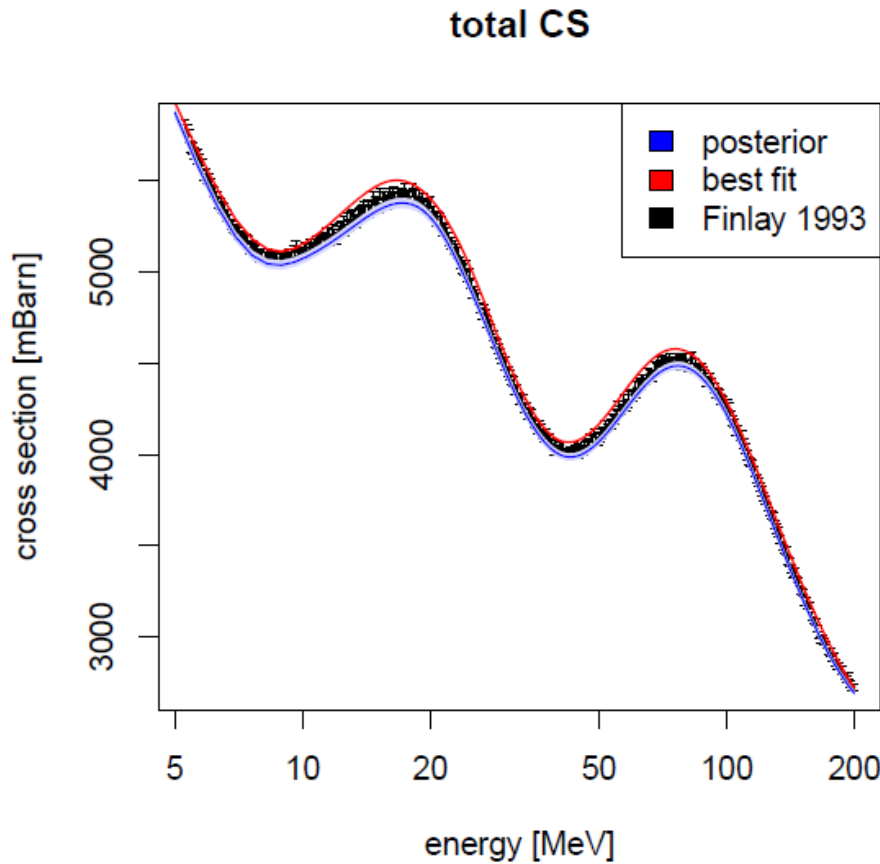
Question:  
 Is this an effect of the well known total cross section?  
 No, the same only with exp. diff. cross sections.





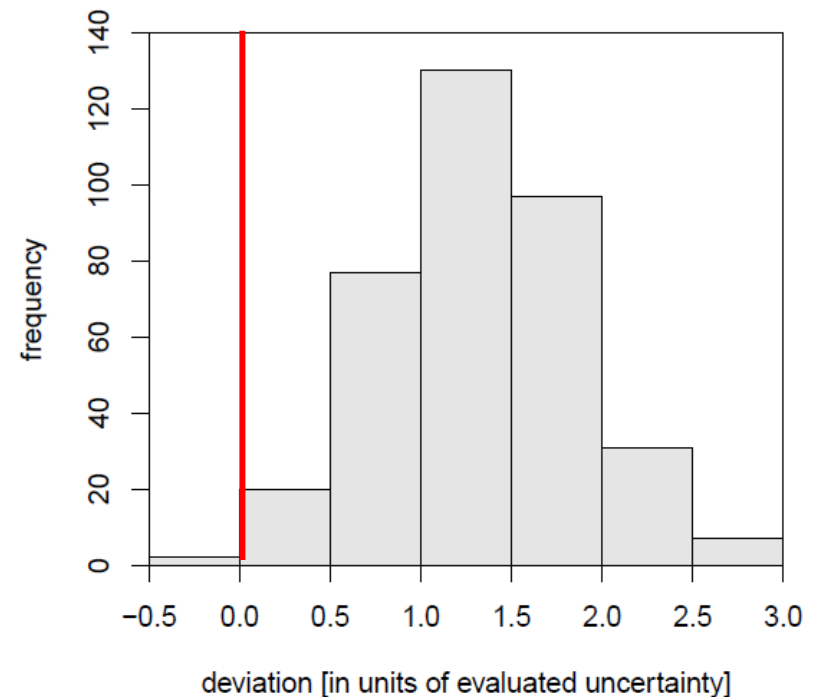
## Evaluation of total cross sections only





- The posterior is systematically lower
- Uncertainties too small

Histogram: deviation of experiment from posterior








Evaluation of total  $n$ - $^{181}\text{Ta}$  cross section only

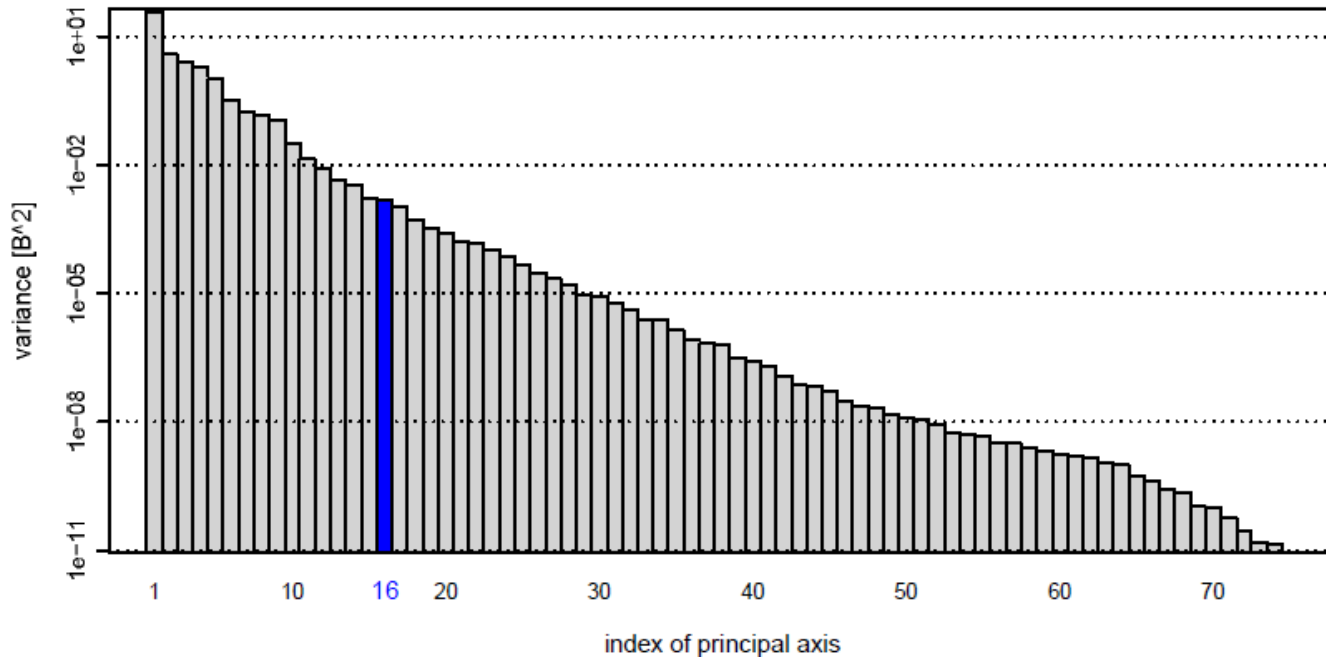
## Outcome of test evaluation

**prior:** accounts for parameter uncertainties only

**data:** n-<sup>181</sup>Ta including either  $\sigma_{\text{tot}}$ ,  $\left(\frac{d\sigma}{d\Omega}\right)_{\text{el}}$  or both  
between 0.3 and 150 MeV

- Inclusion of normalisation necessary 
- Evaluation of  $\sigma_{\text{tot}}$  is too small 
- Uncertainties of  $\sigma_{\text{tot}}$  are too small 
- Evaluation of diff.data systematically too low 
- Uncertainties of evaluated diff.data unrealistically small 

## Eigenvalues and Eigenvectors of covariance matrix $A^{PU}$

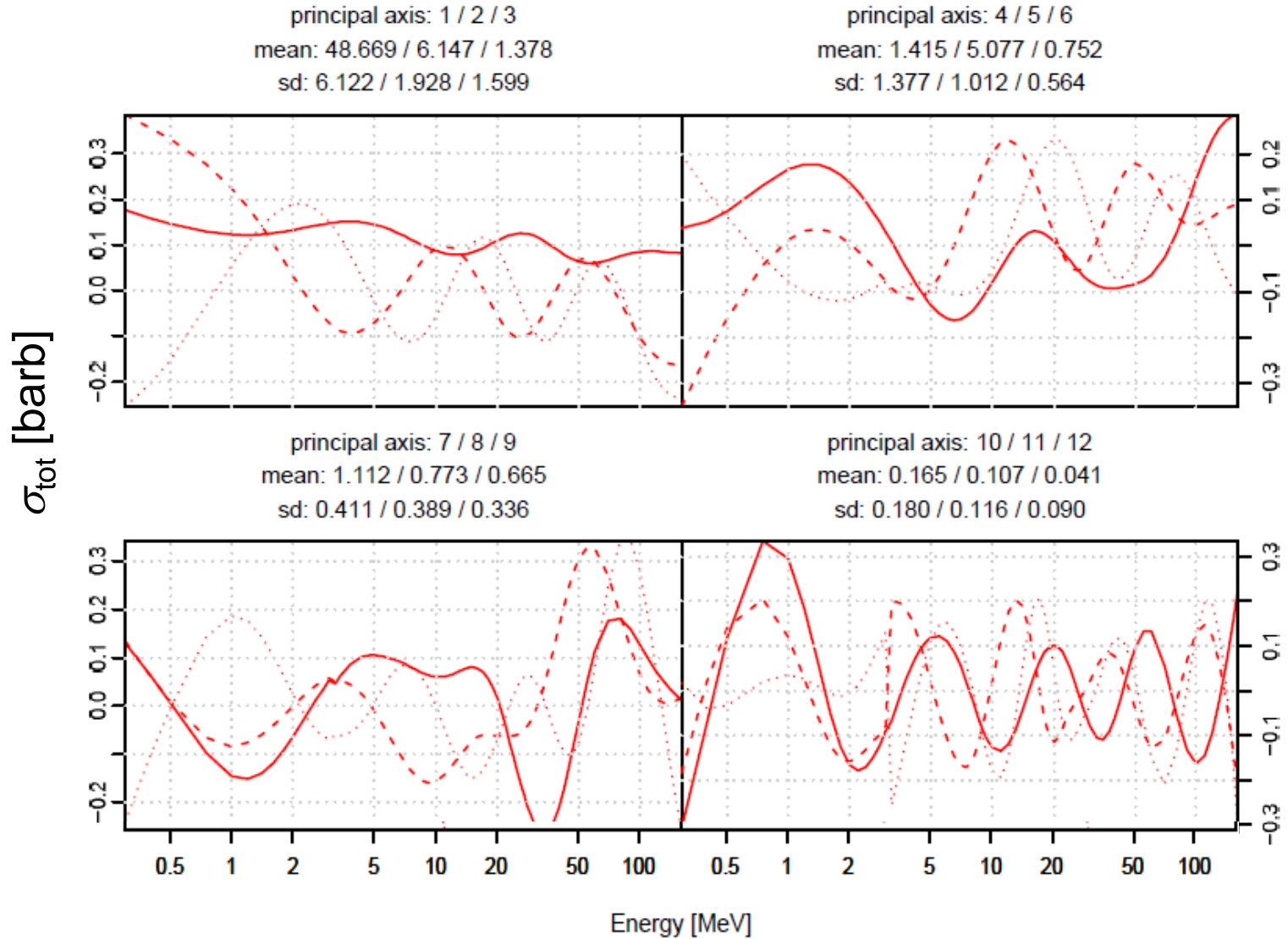


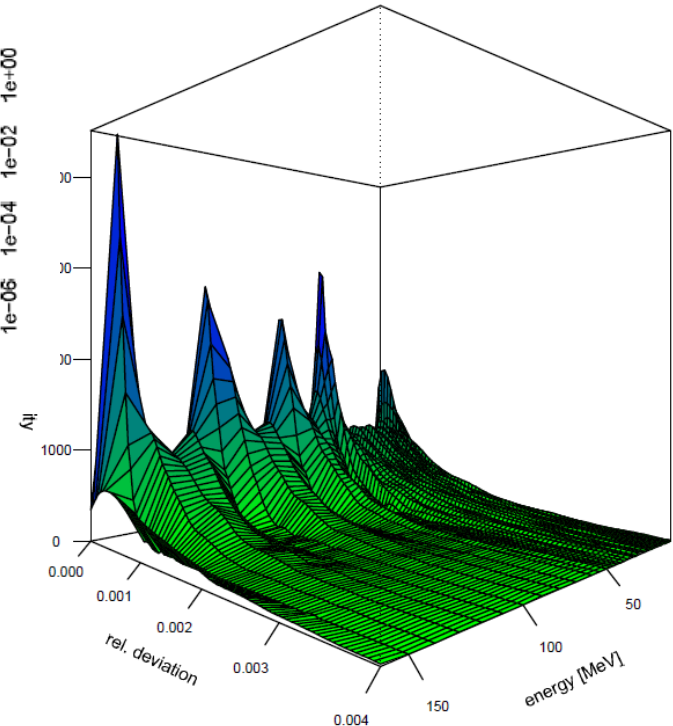
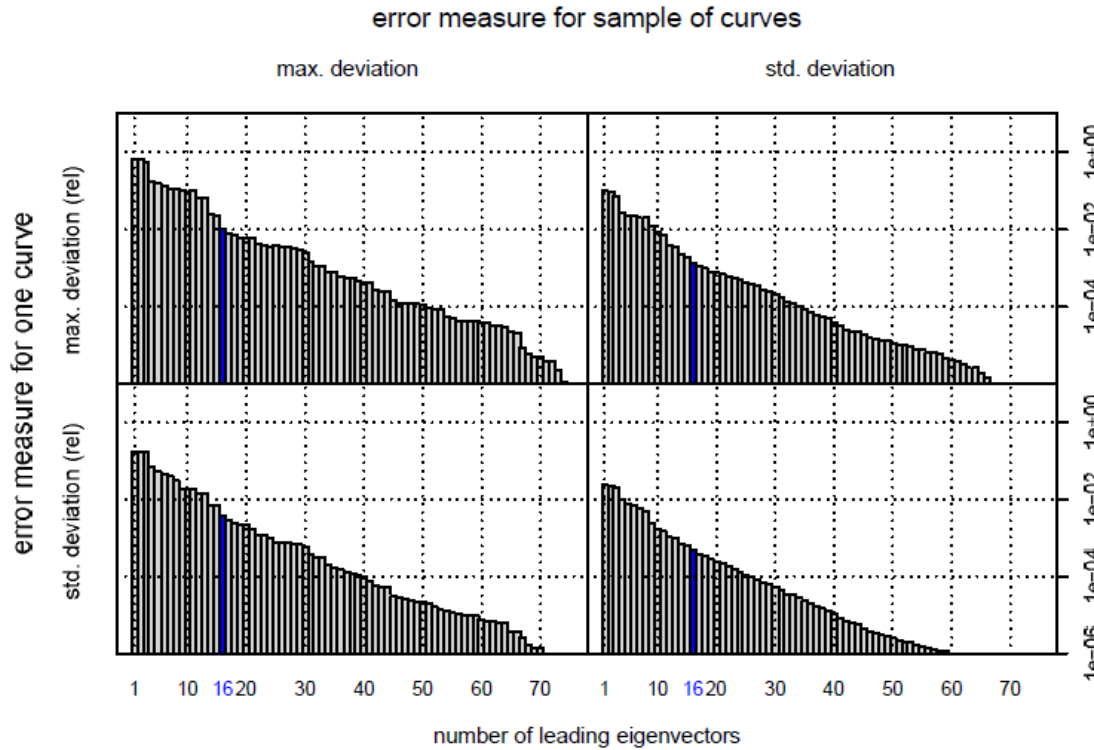
$\sigma_{tot}$  depends on 16 optical model parameters  $\rightarrow$  16 dof of cross section subspace

Eigenvalues = variances in direction of axes

Largest eigenvalue = most significant information

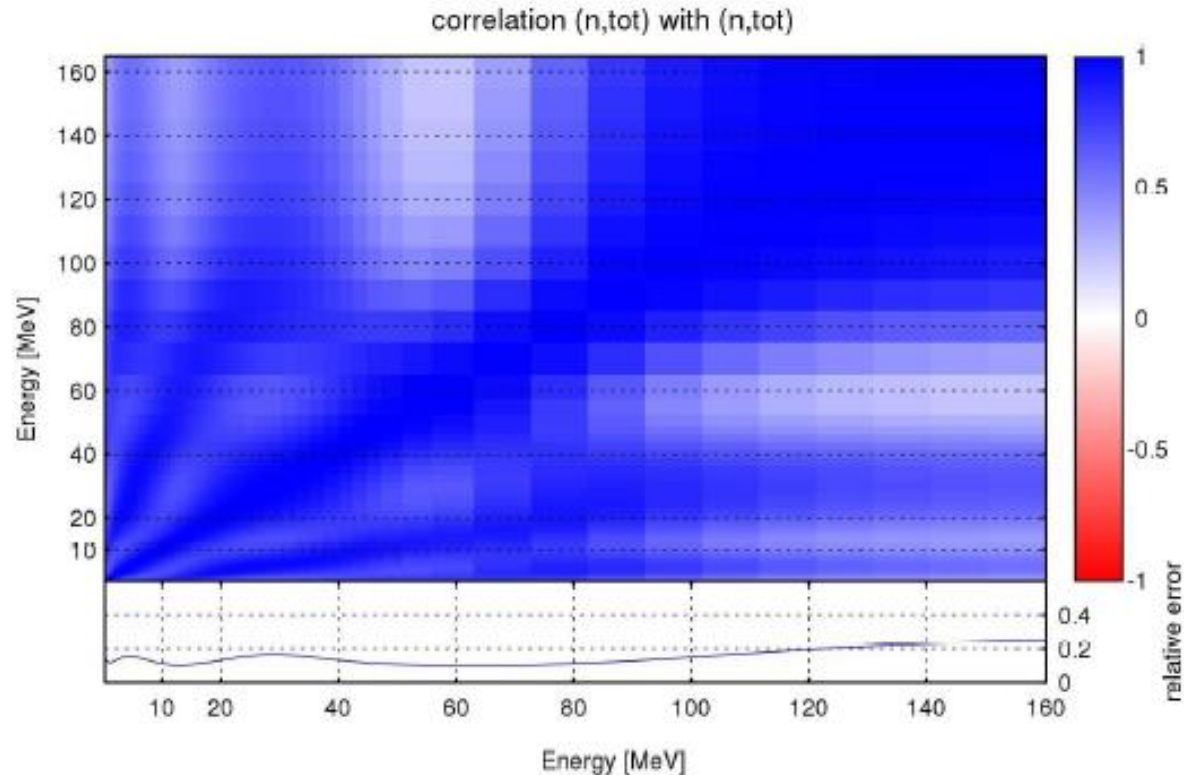
# Principal component analysis





Spectral representation of the covariance matrix by the first 16 eigenvectors leads to accuracy better than 0.3% for  $\sigma_{\text{tot}}(E)$

Restricting to 16 eigenvectors leads to an almost perfect reproduction of the correlation matrix for total xsections

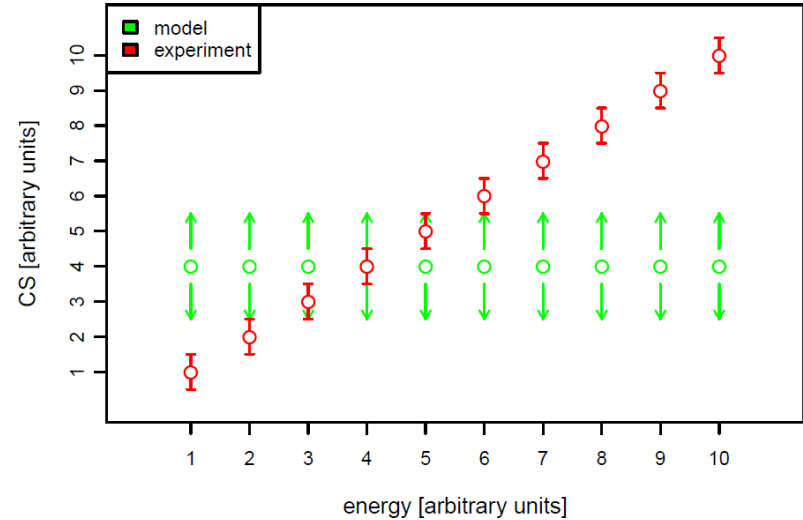




**Problem:** Quality of the model is essential

Experimental Data

Primitive schematic model  
(only one finite eigenvalue)



$$\underline{x}_{post} = \frac{x_{prior} + \bar{y} \frac{n\delta_{prior}^2}{\Delta_{stat}^2 + n\Delta_{syst}^2}}{1 + \frac{n\delta_{prior}^2}{\Delta_{stat}^2 + n\Delta_{syst}^2}} \begin{pmatrix} 1 \\ 1 \\ \vdots \\ 1 \\ 1 \end{pmatrix}$$

$$\bar{y} = \frac{1}{n} \sum_{i=1}^n y_i$$

$$\underline{A}_{post} = \frac{\delta_{prior}^2}{1 + \frac{n\delta_{prior}^2}{\Delta_{stat}^2 + n\Delta_{syst}^2}} \underline{J}_{=n} \quad \text{mit} \quad \underline{J}_{=n} = \frac{1}{n} \begin{pmatrix} 1 & 1 & \dots & 1 & 1 \\ 1 & 1 & \dots & 1 & 1 \\ \vdots & \vdots & \ddots & \vdots & \vdots \\ 1 & 1 & \dots & 1 & 1 \\ 1 & 1 & \dots & 1 & 1 \end{pmatrix}$$

If model strongly correlates, but excludes the shape of the experiments  $\rightarrow$  the posterior becomes arbitrary (bad)

The a-posterior can only be constructed from basis functions available in the covariance matrix

$$p(x | \sigma M) = \frac{p(\sigma | x M) p(x | M)}{p(\sigma | M)}$$



The experimental cross section data for  $^{181}\text{Ta}$  contain only eigenfunctions with very small variances and are contained in the ensemble of cross sections only in a very small fraction

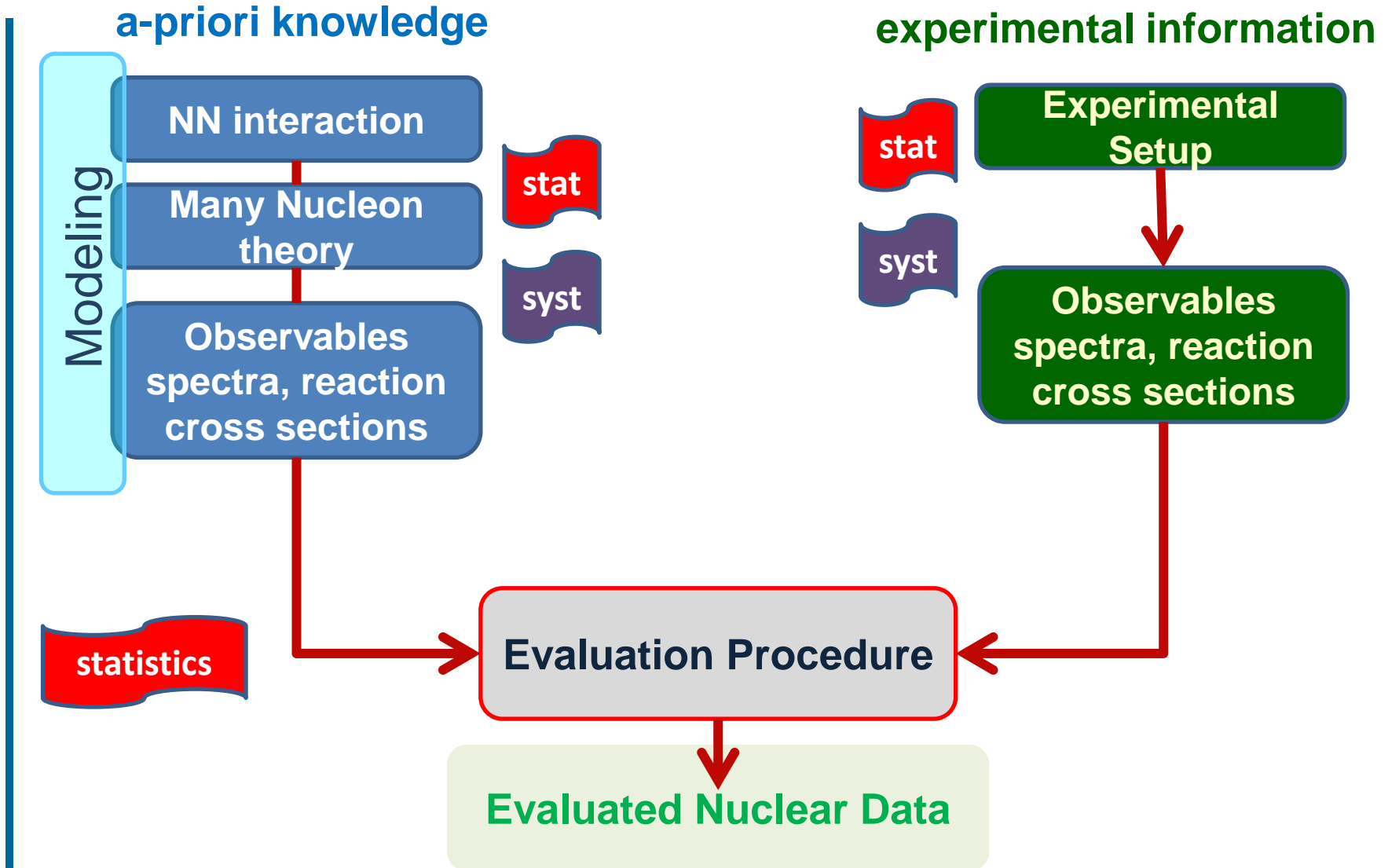


the a-posterior uncertainties are only given by this small fraction by the parameter uncertainties of the nuclear model

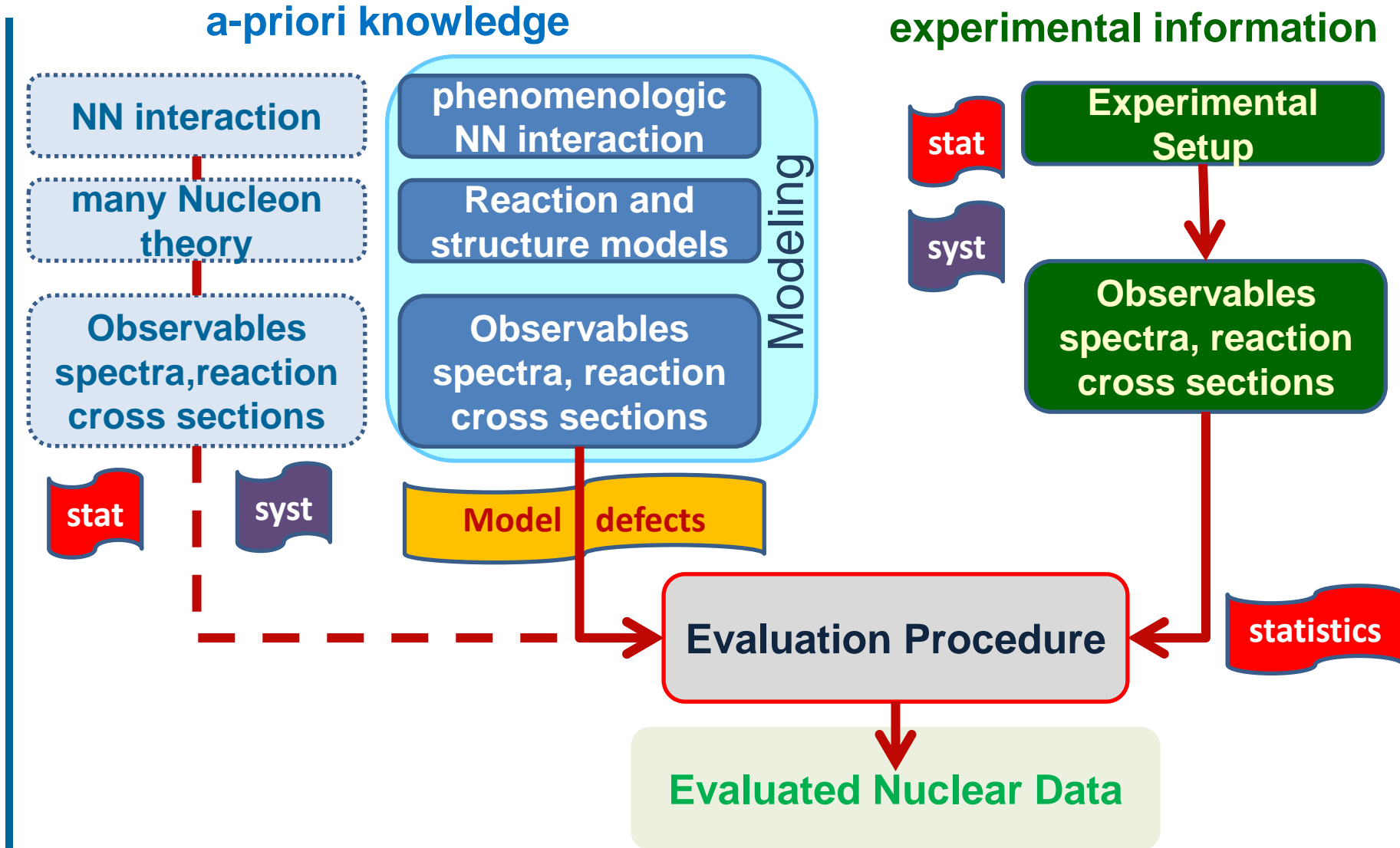


**The strong correlations of the prior must become weaker**

# Concept of Evaluation



# Concept of Evaluation



Systematic deviation of nuclear model values from experimental value which cannot be accounted for within the model by variations of parameters.

D. Neudecker, H. Leeb, R. Capote,  
NIM A (2013)

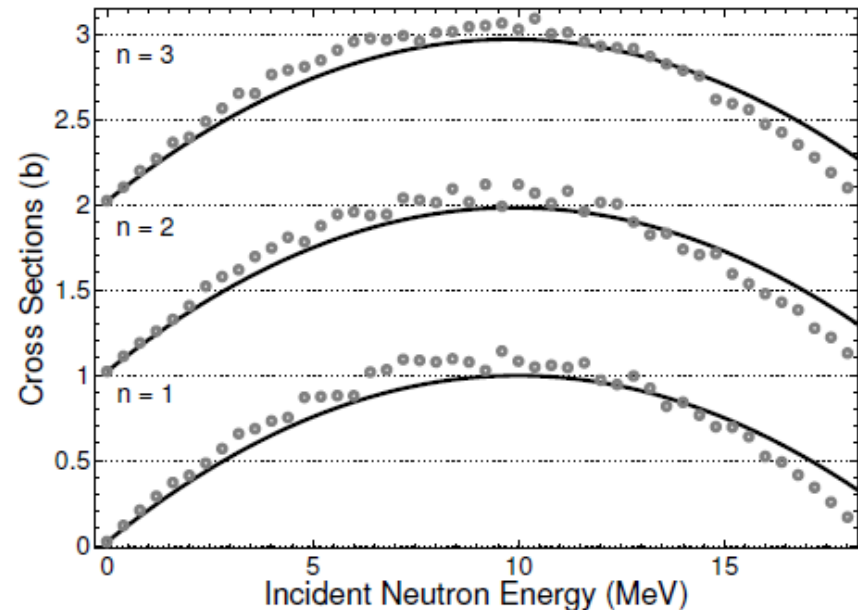


Figure 1: Schematic example of model data (black lines) which deviate systematically from experimental data (gray circles) for  $n = \{1, 2, 3\}$  isotopes above 14 MeV.

# Determination of prior

Nuclear model calculations are used to determine the PRIOR:

The contributions to the covariance matrix of the model are:

$$M^{(mod)} = M^{(par)} + M^{(num)} + M^{(def)}$$

Parameter uncertainties

Deficiencies of the model, is of non-statistical nature

Numerical implementation errors

↑  
neglected

Method and code based on complete ignorance and transformation group invariance developed by Pigni and Leeb

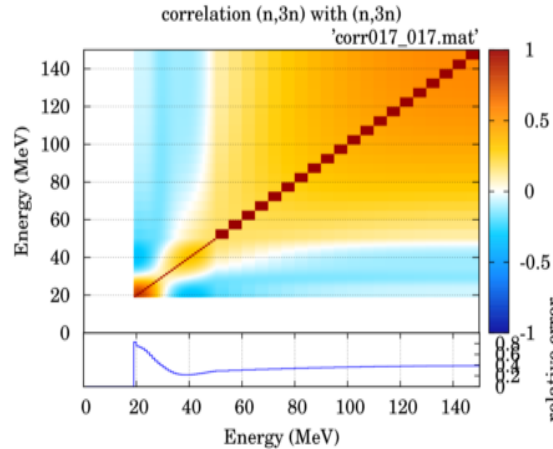
Improved method developed and first application to <sup>55</sup>Mn

H. Leeb et al., Nucl.Data Sheets 109, 2762 (2008)

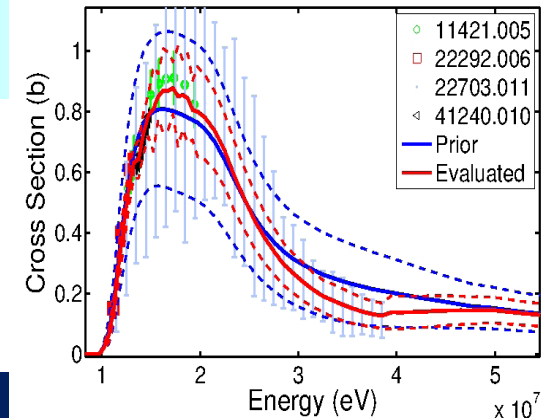
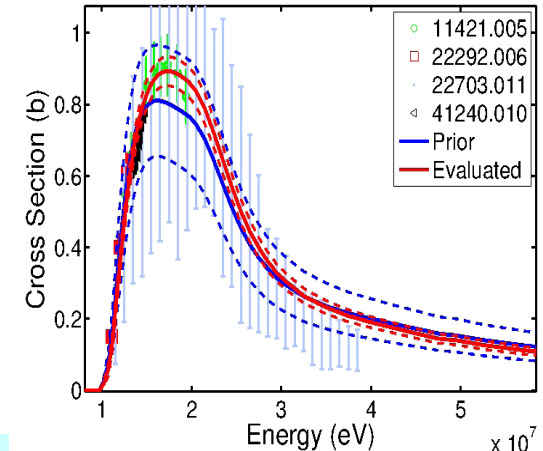
$$D^c = \sum_{n=1}^{N^c} W^{c,n} \sum_{m=1}^M W_m^{c,n} \sum_{j \in Ebin} W_j^{c,n,m} \frac{\sigma_{ex}^{c,n}(E_j)}{\sigma_{th}^{c,n}(E_j)}$$

$$\begin{aligned} \text{cov}(\Delta\sigma^c(E), \Delta\sigma^{c'}(E')) &= A_0^{MD}(E_m, E_{m'}) \\ &= \frac{\sigma_{th}^c(E_m)\sigma_{th}^{c'}(E_{m'})}{\sqrt{N^c(E_m)}\sqrt{N^{c'}(E_{m'})}} \end{aligned}$$

$$\times \left\{ \sum_{n=1}^{N^c} \left[ \langle D_n^c(E_m) \rangle - D^c \right] \left[ \langle D_n^{c'}(E_{m'}) \rangle - D^{c'} \right] + \sum_{n=1}^{N^c} \delta_{cc'} g_{mm'} \sqrt{\text{var}(D_n^c(E_m))\text{var}(D_n^{c'}(E_{m'}))} \right\}$$



without model defects

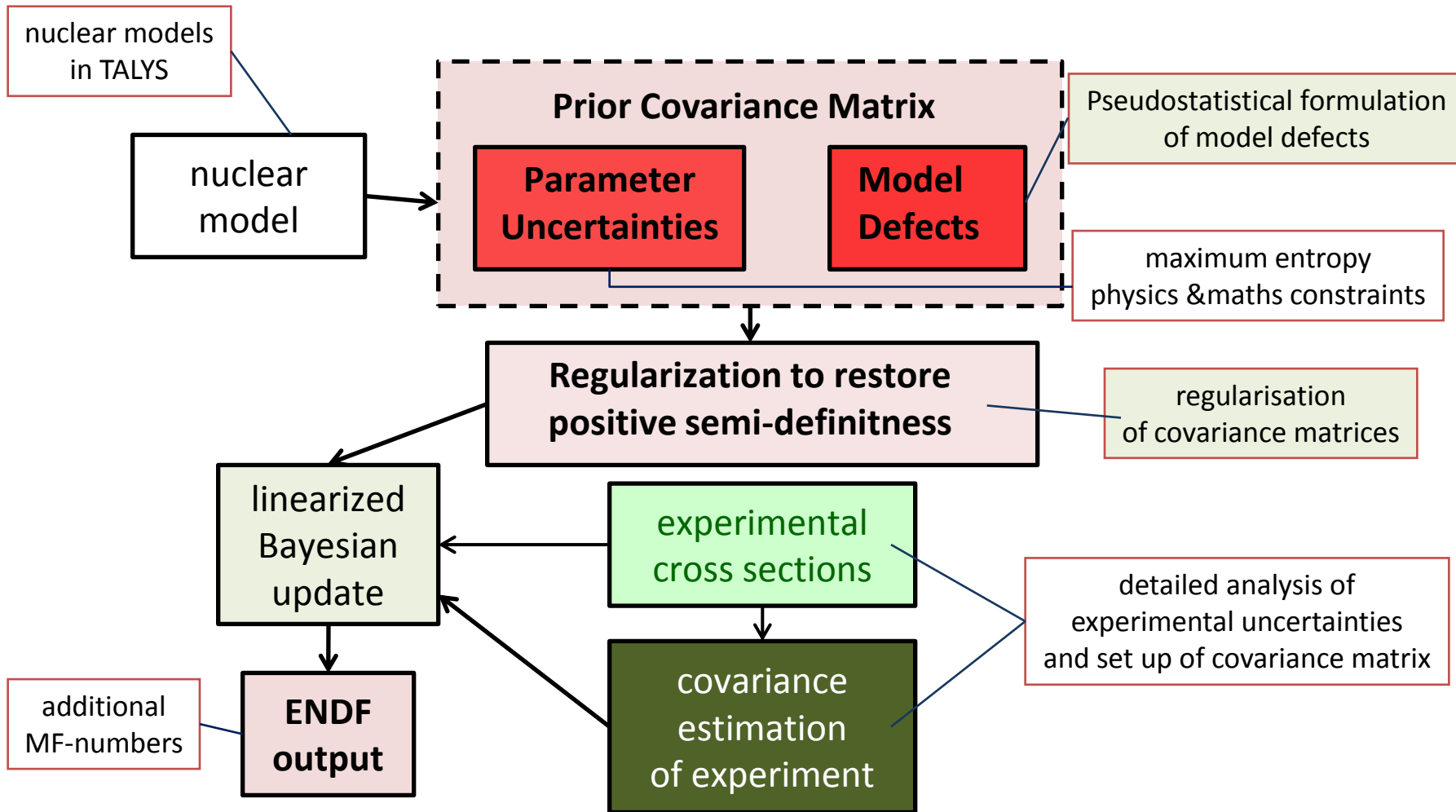


without model defects

Other Procedures for model defects:  
Trkov, R. Capote, Soukhovitsii et al.,  
NIM A212 (2011) 3098

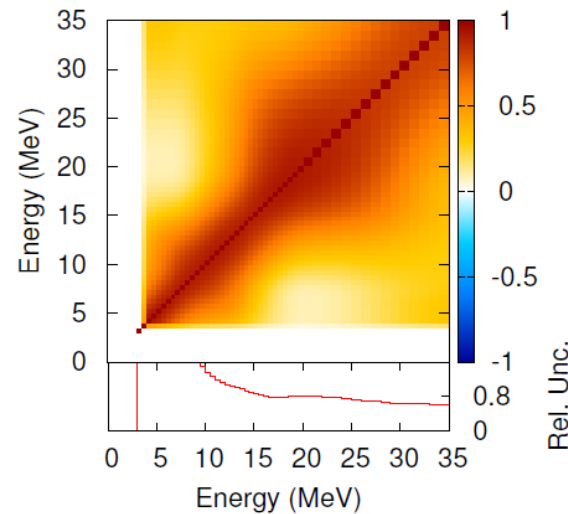
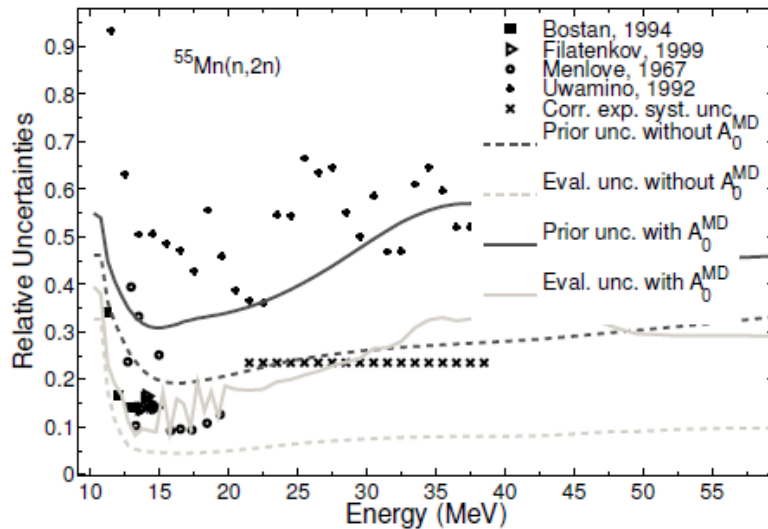
model defects are essential if model deviates significantly from experiment

# Layout of Full Bayesian Evaluation Technique

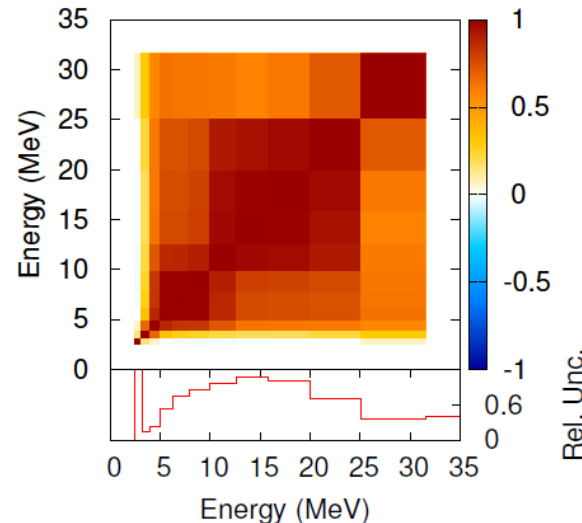




D. Neudecker, H. Leeb, R. Capote,  
NIM A (2013)

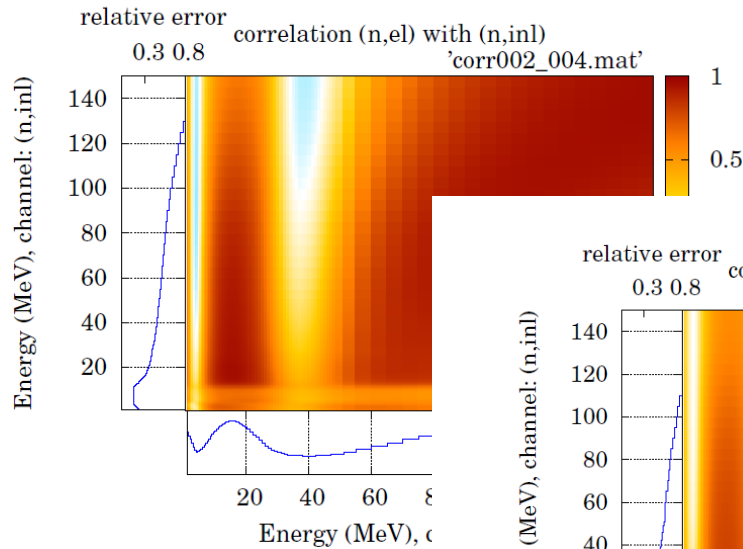


Model defects of Leeb et al., NDS 109 (2008)2762

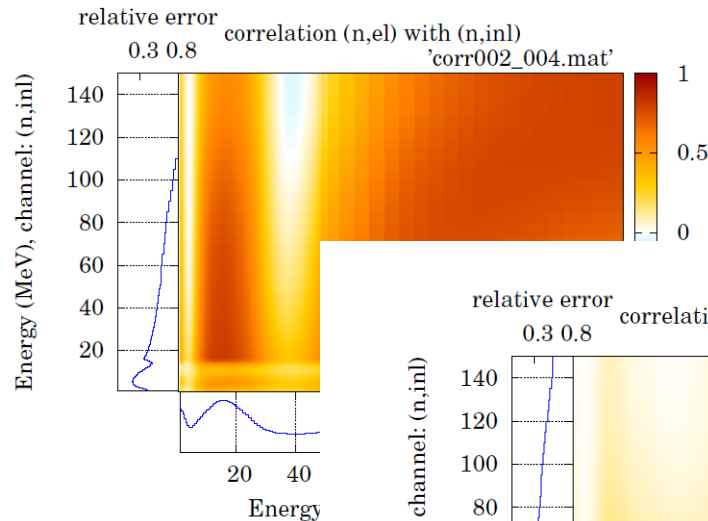


Model Defects of Trkov, Capote et al., NDS 112 (2011)3098

Figure 4: Dimensionless evaluated and prior  $^{55}\text{Mn}(n,2n)$  relative uncertainties including and omitting model defect uncertainties are compared to experimental relative uncertainties of [26–28,30] as well as to the correlated experimental uncertainty above 20 MeV.

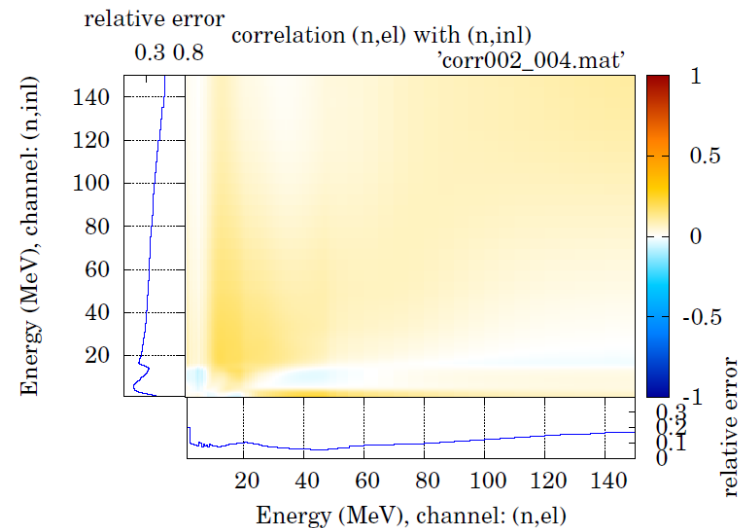


parameter  
uncertainties



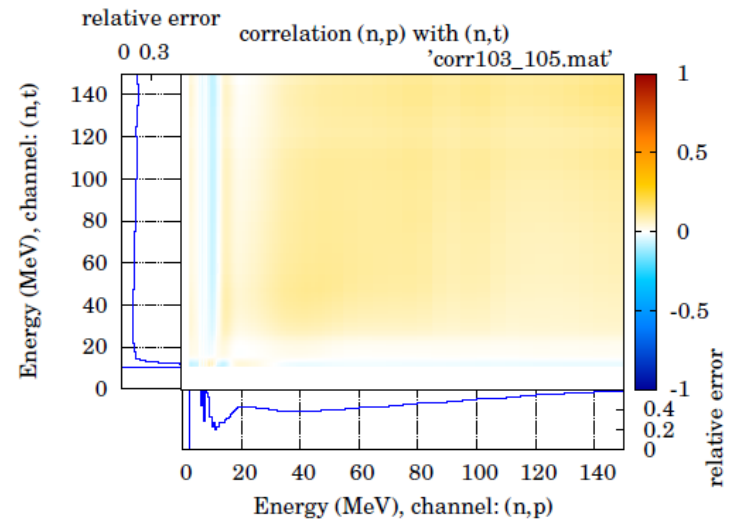
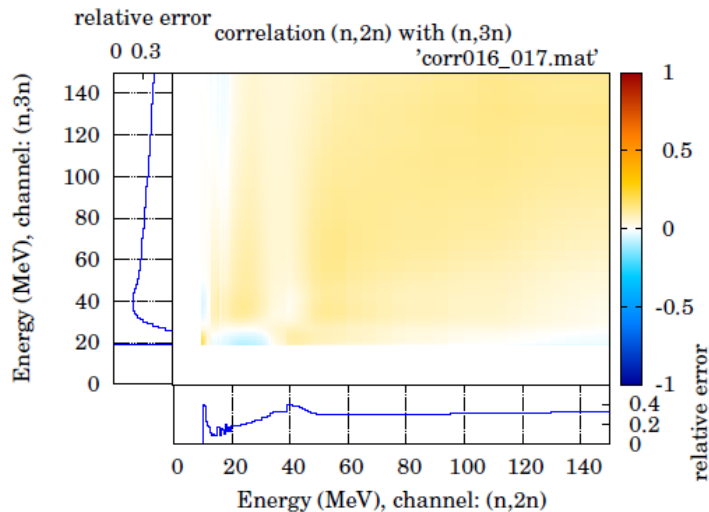
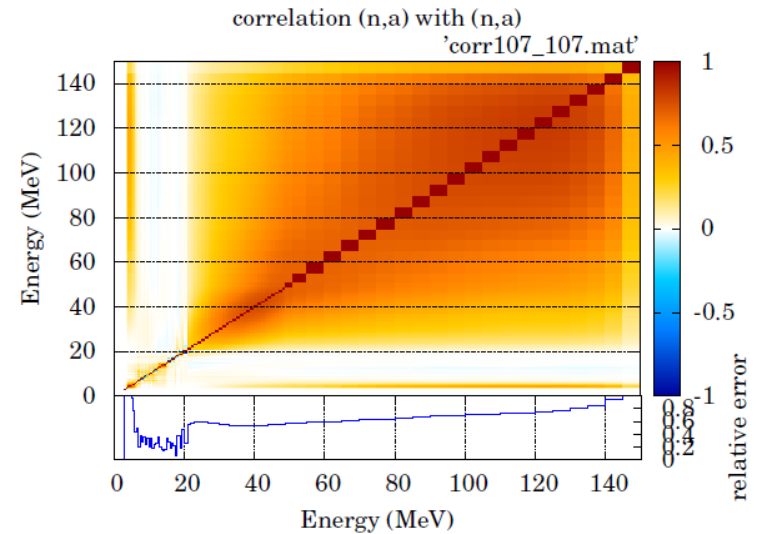
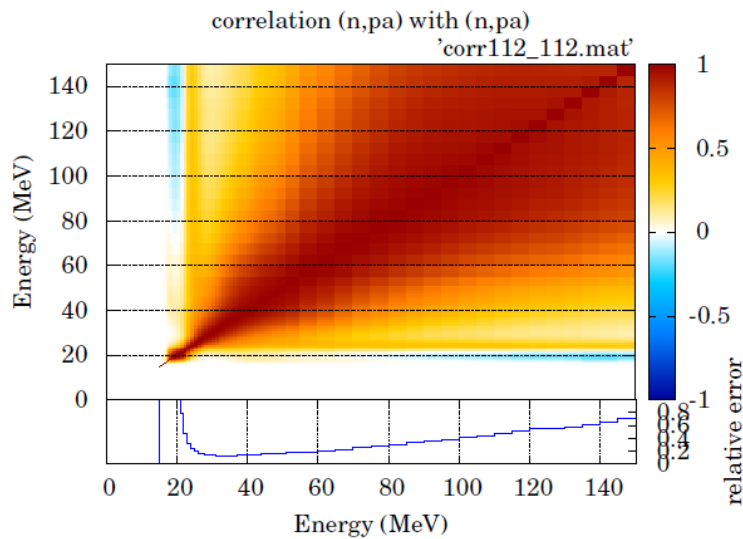
prior

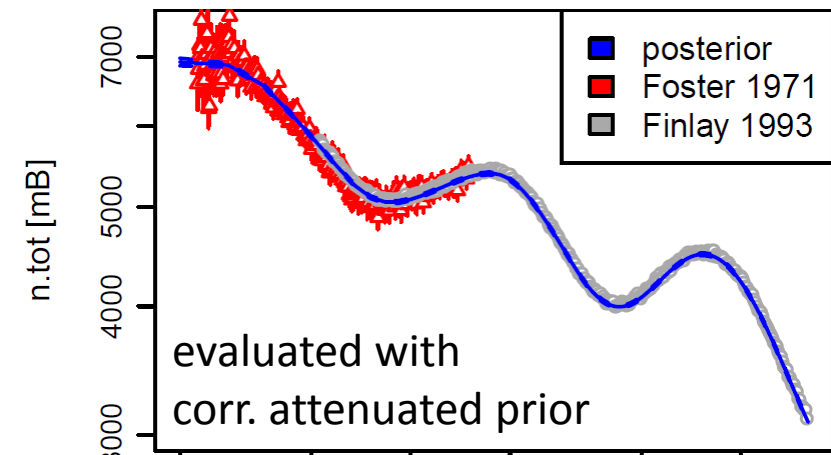
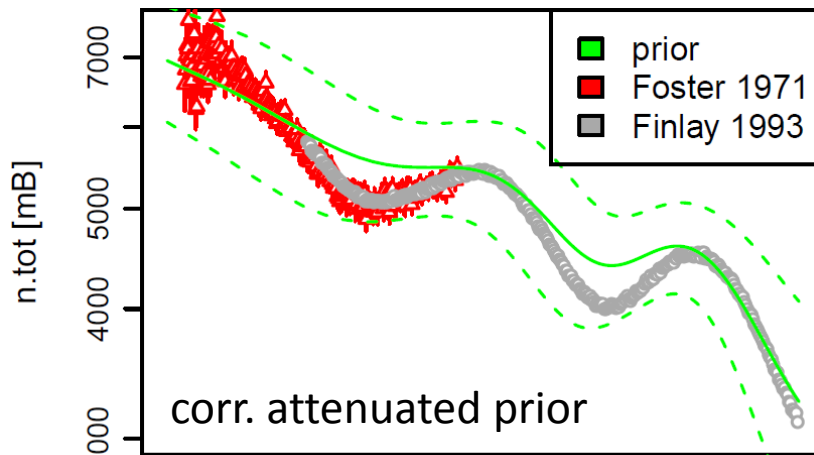
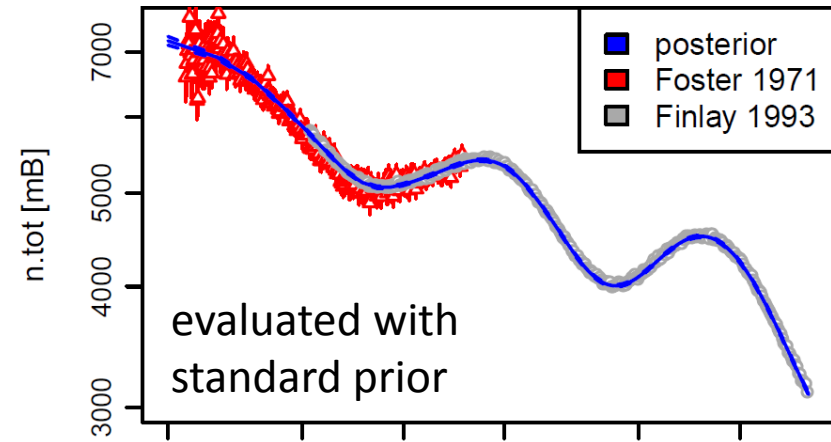
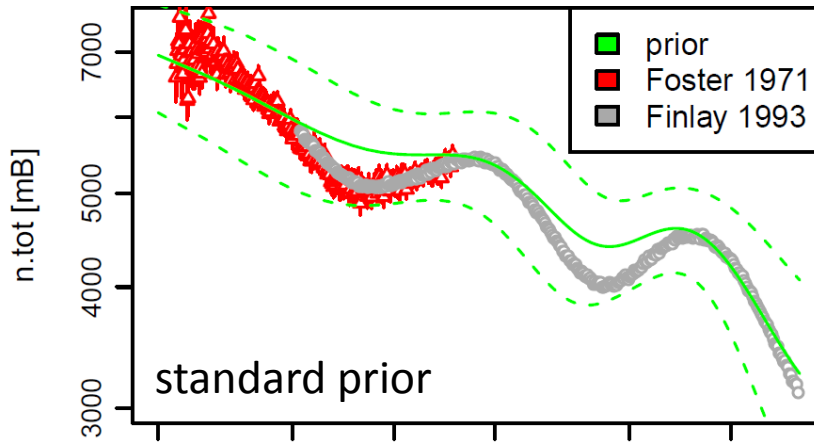
$$\langle \sigma_{\text{ela}}(E) \sigma_{\text{inl}}(E') \rangle$$



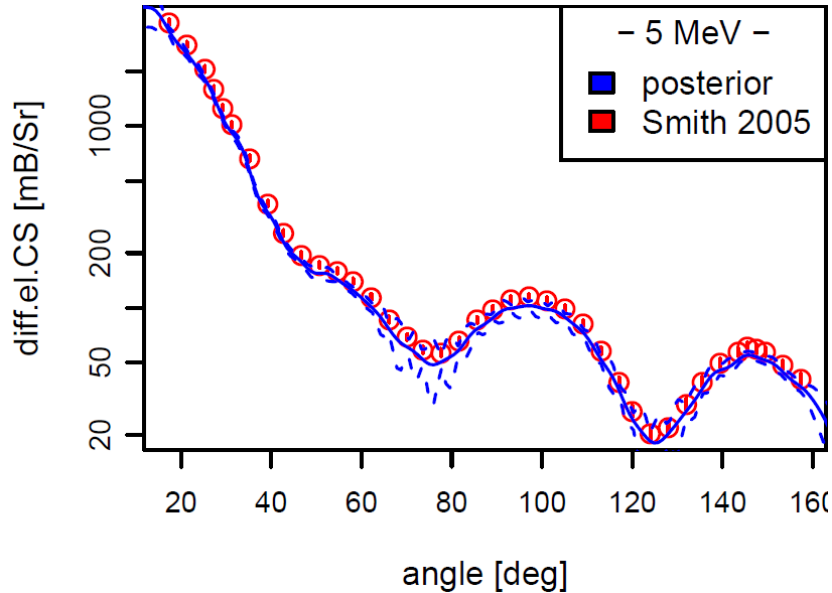
full  
evaluation

<sup>55</sup>Mn

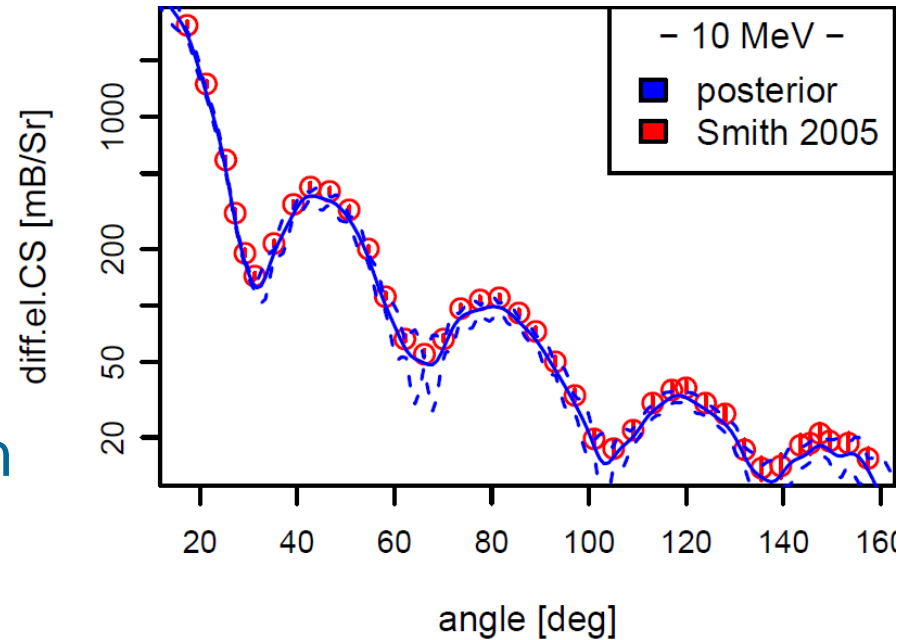




# Example: differential n-<sup>181</sup>Ta elastic cross section



As expected the uncertainties remain in the order of the systematic errors



## Problem:

The pseudostatistical inclusion of model defect covariances violates sum rules, unitarity ...

## Basics of Evaluation Procedures

- The Bayesian evaluation – prerequisite is a perfect model
- Systematic uncertainties – correlated in one step
- Model defects are essential for a realistic evaluation
- Best physics model is required for a proper evaluation

## Future Steps Required

- Inclusion of model defects without violation of sum rules,  
...
- Define meaning of covariance matrices – definition of validation procedure
- Continuous improvement of modeling

# The Hen-Egg Problem

Observable:  $\sigma = \sigma(x, \underline{p})$

parameter  $p_i, i=1, \dots, N$

$$\langle \Delta\sigma(x)\Delta\sigma(x') \rangle = \sum_{i,j=1}^N \frac{\partial\sigma(x)}{\partial p_i} \langle \Delta p_i \Delta p_j \rangle \frac{\partial\sigma(x')}{\partial p_j}$$

observables:

correlated

-----

uncorrelated

parameters:

uncorrelated

-----

correlated



General criterion for covariances required !!!

e.g. for application, comparison

Observable:  $y(x) = f(x, \vec{p}) = g(x, \vec{q})$       $p_i \quad i = 1, \dots, N$

exact description     model

$q_j \quad j = 1, \dots, M$

introduce grid in  $x_i \quad i = 1, 2, \dots, M$       $(\underline{S})_{i,j} = \frac{\partial f(x_i, \underline{p})}{\partial p_j}$

$$\underline{\underline{F}} = \langle\langle \Delta p_i \Delta p_j \rangle\rangle \qquad \underline{\underline{G}} = \langle\langle \Delta q_i \Delta q_j \rangle\rangle$$

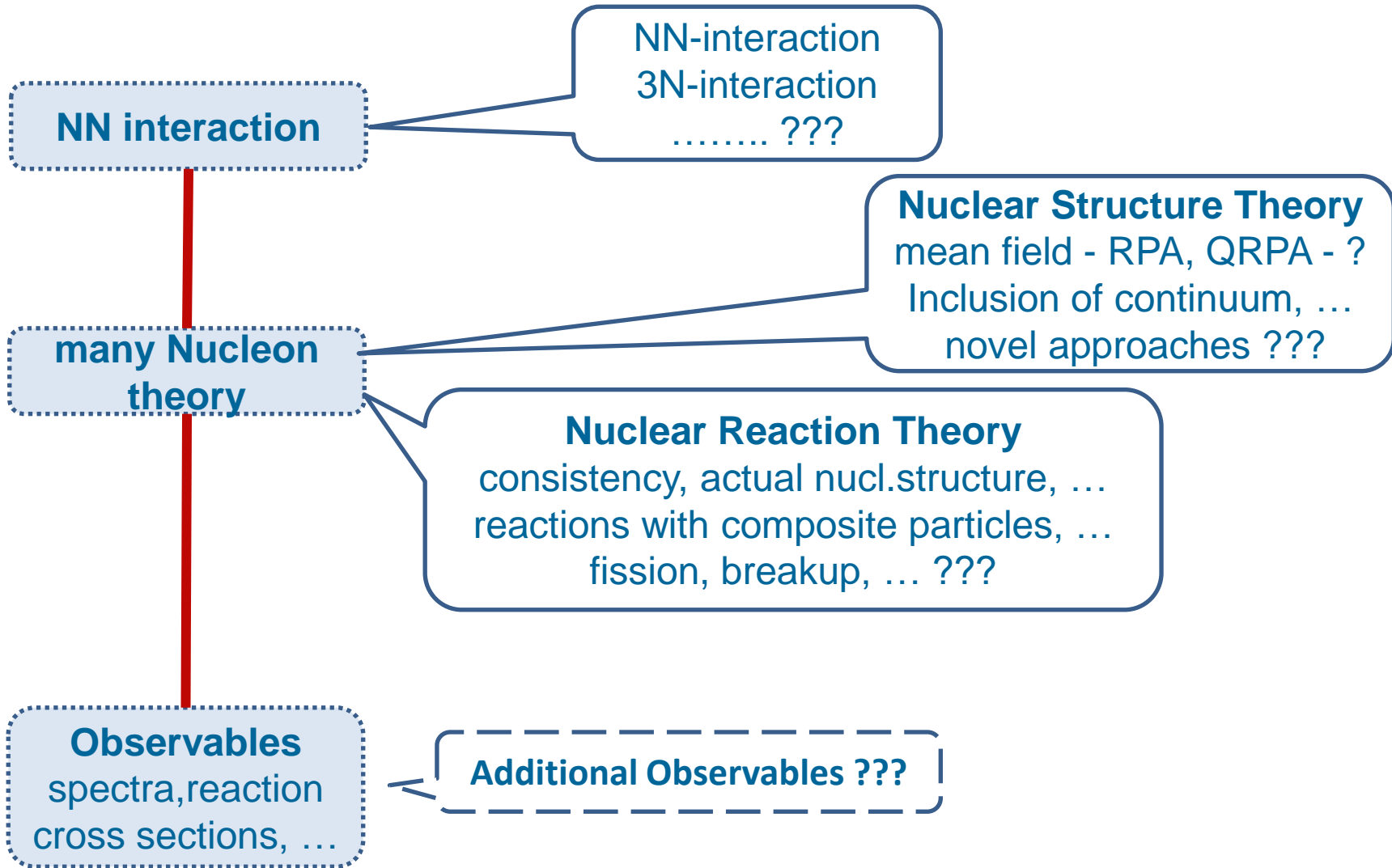
$$\left. \begin{aligned} \langle\langle \Delta y(x_i) \Delta y(x_j) \rangle\rangle^{ph} &= \underline{\underline{Y}}^{ph} = \underline{\underline{S}} \underline{\underline{F}} \underline{\underline{S}}^T \\ \langle\langle \Delta y(x_i) \Delta y(x_j) \rangle\rangle^{mo} &= \underline{\underline{Y}}^{mo} = \underline{\underline{R}} \underline{\underline{G}} \underline{\underline{R}}^T \end{aligned} \right\}$$

Assumption:

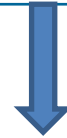
$$\underline{\underline{Y}}^{ph} = \underline{\underline{Y}}^{mo}$$

- If  $M > N$  : for diagonal  $\underline{\underline{F}}$ ,  $\underline{\underline{G}}$  contains dependencies
- If  $M = N$  : corresponds to a rotation in parameter space
- If  $M < N$  : covariance  $\underline{\underline{G}}$  is projection of  $\underline{\underline{F}}$





- Model defects are essential for a realistic evaluation
- Best physics model is required for a proper evaluation



***Nuclear Data Evaluation is a  
Challenge for Nuclear Theory***

**Thank you for your attention**