

Summary of Derivations and Equivalence between Bias Factor Methods and Adjustment Methods

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Action Agreed at June 2019 Meeting*

- **K. Yokoyama** to provide papers to **G. Palmiotti** which illustrate the equivalence of different bias factors methods and their equivalence to the extended adjustment method



- During the June 2019 meeting, eight papers listed in the last slide were delivered personally
- Through the mailing list of SG46, the list of the papers were distributed on September 19, 2019
- In this presentation, I would like to summarize the papers focusing on the equivalence between bias factor methods and adjustment methods

*: Summary Record, Meeting of WPEC SG46, 25-26 June 2019, NEA/NSC/WPEC/DOC(2019)4

Nomenclature in comparison with SG39's

- $P(\mathbf{A}|\mathbf{B})$: conditional probability of \mathbf{A} given \mathbf{B}
- $E(\mathbf{A})$: expectation of \mathbf{A}
- $V(\mathbf{A})$: variance of \mathbf{A}
- \mathbf{T}_0 : unadjusted cross sections (= $\boldsymbol{\sigma}$ in the SG39's common nomenclature)
- \mathbf{T}_x : adjusted cross sections by methodology x (= $\boldsymbol{\sigma}'$)
- $\mathbf{R}_e^{(1)}$: measured value of integral experiments (= \mathbf{E})
- $\mathbf{R}_c^{(1)}(\mathbf{T})$: calculation value of integral experiments (= \mathbf{C})
- $\mathbf{R}_x^{(2)}$: design value of the target system by methodology x
- $\mathbf{G}^{(1)}$: sensitivity matrix of integral experiments (= \mathbf{S})
- $\mathbf{G}^{(2)}$: sensitivity matrix of integral parameters of the target system
- \mathbf{M} : covariance matrix of unadjusted cross sections (= \mathbf{M}_σ)
- $\mathbf{V}_e^{(1)}$: covariance matrix of experimental error for the integral experiments (= \mathbf{M}_E)
- $\mathbf{V}_m^{(1)}$: covariance matrix of analysis method error for the integral experiments (= \mathbf{M}_C)
- $\mathbf{V}_{e+m}^{(1)} = \mathbf{V}_e^{(1)} + \mathbf{V}_m^{(1)}$ (= \mathbf{M}_{EC})
- $\mathbf{V}_m^{(2)}$: covariance matrix of analysis method error for the target system
- $\mathbf{V}_m^{(12)}$: cross-correlation between the integral experiments and the target system
- \mathbf{F}_x : combination factor matrix in the linear estimation by methodology x

CBCA (= Cross-section Adjustment, GLLS, etc.)

- Classical Bayesian Conventional XS Adjustment method

$$T_{\text{CBCA}} \equiv \underset{\hat{T}}{\operatorname{argmax}} P \left(\hat{T} \mid R_e^{(1)} \right)$$

← estimate

$$P \sim \mathcal{N}(\mu, \Sigma) \quad \text{Gaussian distribution}$$



e.g. WPEC/SG33 intermediate report

$$T_{\text{CBCA}} = T_0 + \mathbf{M}\mathbf{G}^{(1)\text{T}} \mathbf{D}^{-1} \left(R_e^{(1)} - R_c^{(1)}(T_0) \right)$$

$$\mathbf{D} \equiv \mathbf{G}^{(1)} \mathbf{M}\mathbf{G}^{(1)\text{T}} + \mathbf{V}_{e+m}^{(1)}$$

- Well-known formula to members of SG46

EBPE & Matrix Form with Approximation

- Extended Bias factor method (Product of Exponentiated values)

$$R_{\text{EBPE}}^{(2)} = R^{(2)}(\mathbf{T}_0) \left(\frac{\prod_i (R_{e,i}^{(1)})^{F_{\text{EBPE},i}}}{\prod_i (R_{c,i}^{(1)})^{F_{\text{EBPE},i}}} \right)$$

$$\approx R^{(2)}(\mathbf{T}_0) \left(1 + F_{\text{EBPE}} \frac{R_e^{(1)} - R_c^{(1)}(\mathbf{T}_0)}{R_c^{(1)}(\mathbf{T}_0)} \right)$$

T. Kugo et al.
JNST 44[12]1509(2007)

K. Yokoyama et al.
JNST 49[12]1165(2012)
Eqs.(51)-(54)

$$F_{\text{EBPE}} \equiv \underset{F}{\operatorname{argmin}} V \left(\frac{\hat{R}^{(2)}}{R_{\text{true}}^{(2)}} \right)$$

- Gaussian distribution is not assumed in EBPE
- EBPE is equivalent to a linear estimation

NB: It is shown that a result of best representativity method by T. Umano et al. is completely equivalent to that of EBPE in Ref.2



$$F_{\text{EBPE}} = \left(\mathbf{G}^{(2)} \mathbf{M} \mathbf{G}^{(1)\text{T}} + \mathbf{V}_m^{(12)\text{T}} \right) \mathbf{D}^{-1}$$

- This formula is similar to CBCA but different

CBEA

- Classical Bayesian Extended XS Ajustment method

$$\mathbf{T}_{\text{CBEA}} \equiv \underset{\hat{\mathbf{T}}}{\operatorname{argmax}} P \left(\hat{\mathbf{R}}^{(2)} \mid \mathbf{R}_e^{(1)} \right)$$

← estimate

$P \sim \mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma})$ Gaussian distribution



$$\mathbf{T}_{\text{CBEA}} = \mathbf{T}_0 + \left(\mathbf{M}\mathbf{G}^{(1)\text{T}} + \mathbf{G}^{(2)} + \mathbf{V}_m^{(12)} \right) \mathbf{D}^{-1} \left(\mathbf{R}_e^{(1)} - \mathbf{R}_c^{(1)}(\mathbf{T}_0) \right)$$

$$\mathbf{T}_{\text{CBEA}} \neq \mathbf{T}_{\text{CBCA}}$$

$$\mathbf{R}_c^{(2)}(\mathbf{T}_{\text{CBEA}}) \approx \mathbf{R}_{\text{EBPE}}^{(2)}$$

K. Yokoyama et al.
JNST 49[12]1165(2012)

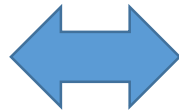
- Design values of CBEA and EBPE are approximately the same

Contrast of Bayesian Inference & Linear Estimation

Bayesian inference

$$P(\hat{\mathbf{T}} | \mathbf{R}_e^{(1)})$$

$$P(\hat{\mathbf{R}}^{(2)} | \mathbf{R}_e^{(1)})$$



$$\mathbf{X}_x \equiv \operatorname{argmax}_{\hat{\mathbf{X}}} P(\hat{\mathbf{X}} | \mathbf{R}_e^{(1)})$$

$$P \sim \mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma})$$

(assumption of Gaussian distribution)

Linear estimation

$$\hat{\mathbf{T}} - \mathbf{T}_0 = \mathbf{F}(\mathbf{R}_e^{(1)} - \mathbf{R}_c^{(1)}(\mathbf{T}_0))$$

$$\hat{\mathbf{R}}^{(2)} - \mathbf{R}_c^{(2)}(\mathbf{T}_0) = \mathbf{F}(\mathbf{R}_e^{(1)} - \mathbf{R}_c^{(1)}(\mathbf{T}_0))$$

$$\mathbf{X}_x = \mathbf{X}_0 + \mathbf{F}_x(\mathbf{R}_e^{(1)} - \mathbf{R}_c^{(1)}(\mathbf{T}_0))$$

$$\mathbf{F}_x \equiv \operatorname{argmin}_F \operatorname{tr}(V(\hat{\mathbf{X}}))$$

$$E(\mathbf{T}_0) = \mathbf{T}_{\text{true}}, \dots$$

(assumption of unbiased estimation)

where $\mathbf{X}_x = \mathbf{T}_x$ or $\mathbf{R}_x^{(2)}$

BFRS (Bias Factor Method by T. Endo et al.)

- Bias Factor Method Using Random Sampling Technique

$$\mathbf{R}_{\text{BFRS}}^{(2)} \equiv \underset{\mathbf{K}}{\operatorname{argmax}} P \left(\hat{\mathbf{R}}^{(2)} \mid \mathbf{R}_e^{(1)} \right)$$

← estimate

$P \sim \mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma})$ Gaussian distribution



$$\mathbf{R}_{\text{BFRS}}^{(2)} = \mathbf{R}_c^{(2)}(T_0) + \mathbf{K} \left(\mathbf{R}_e^{(1)} - \mathbf{R}_c^{(1)}(T_0) \right)$$

$$\mathbf{K} \equiv \left(\underset{\mathbf{G}^{(2)} \mathbf{M} \mathbf{G}^{(1)T}}{\operatorname{Cov} \left(\mathbf{R}_c^{(2)}, \mathbf{R}_c^{(1)} \right) + \mathbf{V}_m^{(12)T}} \right) \left(\underset{\mathbf{G}^{(1)} \mathbf{M} \mathbf{G}^{(1)T}}{\operatorname{Cov} \left(\mathbf{R}_c^{(1)}, \mathbf{R}_c^{(1)} \right) + \mathbf{V}_{e+m}^{(1)}} \right)^{-1}$$

$$\mathbf{R}_{\text{BFRS}}^{(2)} \approx \mathbf{R}_{\text{EBPE}}^{(2)}$$

T. Endo et al.
JNST 53[10]1494(2016)

- Design values of BFRS and EBPE are approximately the same

MLEA

- MVULE (Minimum Variance Unbiased Linear Estimation)-based Extended XS Adjustment method

$$\mathbf{T}_{\text{MLEA}} = \mathbf{T}_0 + \mathbf{F}_{\text{MLEA}} \left(\mathbf{R}_e^{(1)} - \mathbf{R}_c^{(1)}(\mathbf{T}_0) \right)$$

$$\mathbf{F}_{\text{MLEA}} \equiv \underset{\mathbf{F}}{\operatorname{argmin}} \operatorname{tr} \left(V \left(\mathbf{R}_c^{(2)}(\hat{\mathbf{T}}) \right) \right) \cap \underset{\mathbf{F}}{\operatorname{argmin}} \operatorname{tr} \left(V(\hat{\mathbf{T}}) \right)$$



$$\mathbf{T}_{\text{MLEA}} = \mathbf{T}_{\text{CBEA}}$$

$$\mathbf{R}_c^{(2)}(\mathbf{T}_{\text{MLEA}}) = \mathbf{R}_c^{(2)}(\mathbf{T}_{\text{CBEA}}) \approx \mathbf{R}_{\text{EBPE}}^{(2)}$$

K. Yokoyama et al.
JNST 56[1]87(2019)

- EA (Extended Adjustment method) can be derived by a linear estimation

MLCA

- MVULE (Minimum Variance Unbiased Linear Estimation)-based Conventional XS Adjustment method

$$\mathbf{T}_{\text{MLCA}} = \mathbf{T}_0 + \mathbf{F}_{\text{MLCA}} \left(\mathbf{R}_e^{(1)} - \mathbf{R}_c^{(1)}(\mathbf{T}_0) \right)$$

$$\mathbf{F}_{\text{MLCA}} \equiv \underset{\mathbf{F}}{\text{argmin}} \text{tr} \left(V(\hat{\mathbf{T}}) \right)$$



$$\mathbf{T}_{\text{MLCA}} = \mathbf{T}_{\text{CBCA}}$$

K. Yokoyama et al.
JNST 56[1]87(2019)

- The conventional cross-section adjustment method can be derived by a linear estimation
- Gaussian distribution is not assumed in MLCA although the assumption of unbiased estimation is required

Concluding Remarks

- The extended bias factor method (EBPE) is approximately equivalent to the extended adjustment method (CBEA) although EBPE does not adjust cross sections explicitly
 - EBPE is derived by a linear estimation
 - CBEA is derived by a Bayesian inference
- The (classical Bayesian) conventional cross-section adjustment method (CBCA) is different from CBEA
 - CBCA maximizes the probability of cross sections
 - CBEA maximizes the probability of design values
- The conventional and the extended cross-section adjustment methods can be derived by a linear estimation (MLCA and MLEA)
 - MLCA minimizes the variance of cross sections
 - MLEA minimizes the variance of design values

References (Provided Papers)

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